Module 14: Advanced Heritability Analysis

- 1 Estimating SNP Heritability
- 2 Genome Partitioning
- Intensity of Heritability
 - 4 Testing Different Models
- Using Other Datatypes
- 6 Bivariate Analysis
 - 7 Mixed Model Analysis
 - Gene-Based Association Testing
 - Adaptive MultiBLUP
- 10 Applying to Animal and Plant Data

Estimating SNP Heritability

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We looked at estimating heritability from SNP data. To do this in a mixed model framework, it is necessary to construct a kinship matrix \mathbf{K} , which estimates relatedness / genetic similarity between all pairs of individuals

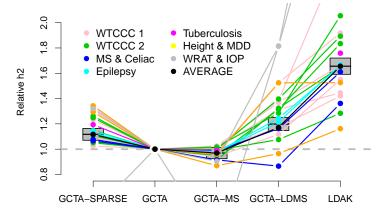
Most commonly, **K** takes the form of allelic correlations $\mathbf{K} = \mathbf{X}\mathbf{X}^{T}$, in which case, the mixed model is equivalent to a random effect regression model

When individuals are related, we will obtain an estimate of narrow-sense heritability, h^2

When individuals are unrelated, (and there is no inflation due to population structure or genotyping error) we will obtain an estimate of SNP heritability, h_{SNP}^2

Recap of Module 9

The standard K used when estimating SNP heritability, corresponds to a (very) specific set of model assumptions



Changing **K** changes estimates of h_{SNP}^2 ; sometimes by a large amount

Different Methods Give Different Estimates

ANALYSIS



A

Improved Heritability Estimation from Genome-wide SNPs

Estimation of narrow-sense heritability, h², from genome-wide SNPs genotyped in unrelated individuals has recently att

Doug Speed,1,* Gibran Hemani,2 Michael R. Johnson,3 and David J. Balding1

Common SNPs explain a large proportion of the heritability for human height nature

genetics

Jian Yang¹, Beben Benyamin¹, Brian I Pamela A Madden², Andrew C Heath² Peter M Visscher¹

SNPs discovered by penome-wide asso account for only a small fraction of the ge an height explained by 294,831 5N 3,925 unrelated individuals using a lines the observed cenotype data. We show fi can be explained by considering all SNPs nost of the heritability is not missing but to pass stringent significance tests. We r that the remaining heritability is due to

disequilibrium between causal variants and exacerbated by causal variants having low

ency than the SNPs canford to dat

Estimation of SN Heritability from Dense Genotype

To the Editor: Recently, Spee hensive and elegant evaluat underlying the linear mixed program GCTA2 for estimation

They concluded that the method is robust to violations of four of the assumptions. However, they found that SNP-heritability estimates were sensitive to uneven linkage disequilibrium (LD) between SNPs (implying uneven tagging of causal variants) and suggested an approach to improving the robustness of estimates in this context. Speed et al.1 tested their method on relatively sparse geno-

Genetic variance estimation with imputed variants finds negligible missing heritability for human height and body mass index

Jian Yang^{1,2,24} arew Jakshi1, Matthew R Ro nson¹, John R I The LifeLines ohort Study⁸. nusson¹⁵, Nan Patrik K E Ma Michael E Coddard² Naomi R Wray

hitecture used in the construction of the GRM They also noted that uneven tagging of causal variants by genotyped SNPs generated biased estimates of here under some genetic architectures. They proposed that SNP contributions should be weighted by the LD (r^2) between SNPs. However, we found that the weighted GRM can generate upwardly biased estimates of h_{SNP}^2 in the context of dense sanotuning because the density distribution of MAE which

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A E Vinkhuyzen Sang Hong Lee^{1,4}, ', Gib. n He anı ..., Ann ana 1:1-Migi⁹ mi⁹. 1 16,17, Nicole Soran ngelss issch 1,2,24

van Vliet-ustaptchouk6,7, He old Snieder6, 18.19

es Metspal 2^{,13}, Anders Hamsten¹⁴, atthew C Keller^{20,21},

and offers several advantages over traditional pedigree-based methods. With the use of this approach, it has been estimated renotyping array. In compari ulation the validity of sever found that the method is rea um (LD) between SNPs: cont ons of low LD. The overall d bstantial in realistic scenario hat this correction greatly re he first seven diseases studie e for immune-related diseas immune diseases. To calcula

> AIHG PMCID

ed Heritability Analysis

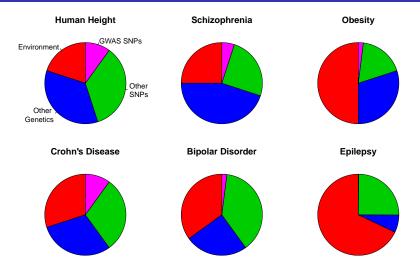
inson.3 and David J. Balding

This article has been cited by other articles in PMC

Main Text

To the Editor: In Speed et al., 1 we identified two potential issues when r SNP-based heritability estimation: (1) estimates of h^2 can be biased whe

The Missing Heritability Problem is Solved



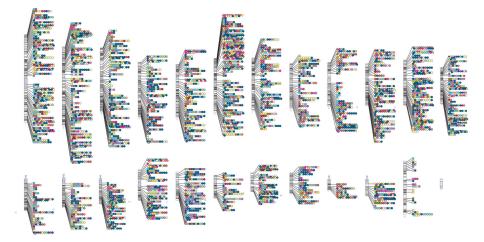
Regardless of the exact numbers, SNPs explain considerable heritability

The three main objectives are:

- 1 Identify the causal variants
- 2 Create prediction models
- 3 Understand the genetic architecture of the trait

The most popular analysis is single-SNP testing

1 - Identify the Causal Variants



Successful for traits with very large causal variants (e.g., BRCA1 for breast cancer, HLA for celiac)

But limited success for most common traits

Would like to know how much causal variation is due to:

- common variation
- rare variation
- additivity
- dominance / recessiveness / epistasis
- coding regions
- copy number changes, frameshifts, splicings, methlation, etc.

Single-SNP GWAS analysis mainly tells us that there are few common variants of strong effect

SNP-Based Heritability Analysis to the Rescue



2 Genome Partitioning

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Genome Partitioning

The basic (single-kinship) model assumes

$$\begin{split} Y &= \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3 + \beta_4 X_4 + \beta_5 X_5 + \beta_6 X_6 + \beta_7 X_7 \\ &+ \beta_8 X_8 + \beta_9 X_9 + \beta_{10} X_{10} + \beta_{11} X_{11} + \beta_{12} X_{12} + \beta_{13} X_{13} + \beta_{14} X_{14} \\ &+ \beta_{15} X_{15} + \beta_{16} X_{16} + \beta_{17} X_{17} + \beta_{18} X_{18} + \beta_{19} X_{19} + \beta_{20} X_{20} + \beta_{21} X_{21} \\ &+ \beta_{22} X_{22} + \beta_{23} X_{23} + \beta_{24} X_{24} + \beta_{25} X_{25} + \beta_{26} X_{26} + \beta_{27} X_{27} + \beta_{28} X_{28} \\ &+ \ldots + \beta_{500\ 000} X_{500\ 000} \\ &+ e. \end{split}$$

Assume
$$\beta_j \sim \mathbb{N}(0, \sigma_g^2/N)$$
 and $e \sim \mathbb{N}(0, \sigma_e^2)$.
Then $Y \sim \mathbb{N}(0, K\sigma_g^2 + I\sigma_e^2)$, where $K = \frac{XX^T}{N}$

We can extend this to

 $Y = \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3 + \beta_4 X_4 + \beta_5 X_5 + \beta_6 X_6 + \beta_7 X_7$ $+ \beta_8 X_8 + \beta_9 X_9 + \beta_{10} X_{10} + \beta_{11} X_{11} + \beta_{12} X_{12} + \beta_{13} X_{13} + \beta_{14} X_{14}$ $+ \beta_{15} X_{15} + \beta_{16} X_{16} + \beta_{17} X_{17} + \beta_{18} X_{18} + \beta_{19} X_{19} + \beta_{20} X_{20} + \beta_{21} X_{21}$ $+ \beta_{22} X_{22} + \beta_{23} X_{23} + \beta_{24} X_{24} + \beta_{25} X_{25} + \beta_{26} X_{26} + \beta_{27} X_{27} + \beta_{28} X_{28}$ $+ \dots + \beta_{500\ 000} X_{500\ 000}$ + e.

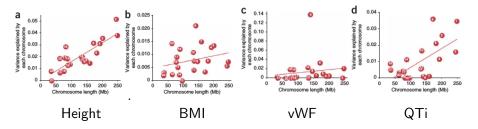
Assume $\beta_j \sim \mathbb{N}(0, \sigma_{g1}^2/N_1)$ and $\beta_k \sim \mathbb{N}(0, \sigma_{g2}^2/N_2)$.

Then $Y \sim \mathbb{N}(0, K_1 \sigma_{g1}^2 + K_2 \sigma_{g2}^2 + I \sigma_e^2)$, where $K_1 = \frac{X_1 X_1^T}{N_1}$ and $K_2 = \frac{X_2 X_2^T}{N_2}$

Instead of estimating only total h_{SNP}^2 , we can now estimate h^2 of RED SNPs and h^2 of BLUE SNPs

Genome partitioning of genetic variation for complex traits using common SNPs

Jian Yang^{1*}, Teri A Manolio², Louis R Pasquale³, Eric Boerwinkle⁴, Neil Caporaso⁵, Julie M Cunningham⁶, Mariza de Andrade⁷, Bjarke Feenstra⁸, Eleanor Feingold⁹, M Geoffrey Hayes¹⁰, William G Hill¹¹, Maria Teresa Landi¹², Alvaro Alonso¹³, Guillaume Lettre¹⁴, Peng Lin¹⁵, Hua Ling¹⁶, William Lowe¹⁷, Rasika A Mathias¹⁸, Mads Melbye⁸, Elizabeth Pugh¹⁶, Marilyn C Cornelis¹⁹, Bruce S Weir²⁰, Michael E Goddard^{21,22} & Peter M Visscher¹





- 2) Genome Partitioning
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Intensity of Heritability

Suppose you have used genome partitioning to estimate contributions of genic and inter-genic regions

$$h_{SNP}^{2} = 60\%$$

$$h_{GENIC}^{2} = 40\%$$

$$h_{INTER-GENIC}^{2} = 20\%$$
Non-Genes

Want to decide whether a partitioning is significant

Define a region's intensity of heritability, I, as its h^2 divided by how much variation it captures (its sum of SNP weightings)

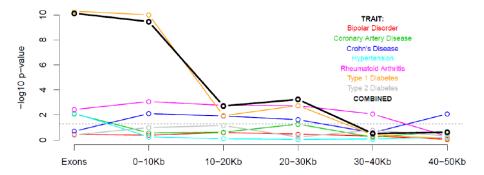
Can then test whether two (or more) partitions have significantly different intensity of heritability

Is intensity of heritability higher for exonic than inter-genic SNPs?

		Intensity of heritability (h ² /1000 "SNPs")		
Trait	Total h ²	Exons	Intergenic	Р
Bipolar Disorder	68%	1.7	1.3	0.37
Coronary Artery Disease	44%	3.1	0.6	0.008
Crohn's Disease	62%	1.6	0.7	0.21
Hypertension	54%	3.6	1.1	0.007
Rheumatoid Arthritis	52%	3.1	0.3	0.004
Type 1 Diabetes	76%	7.5	0.3	5e-11
Type 2 Diabetes	47%	0.9	0.6	0.40

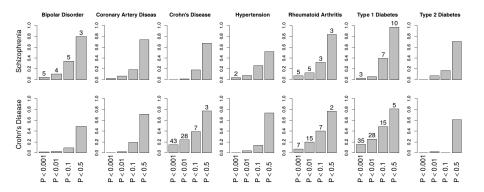
Inter-genic defined as >100kb from a coding region.

The Role of Genes

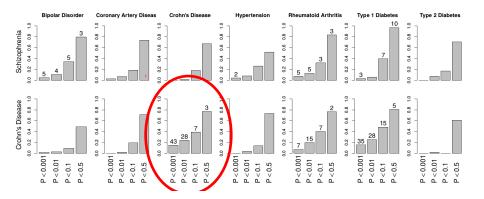


Can investigate what happens as we move away from exons.

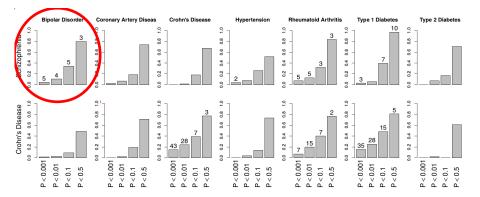
Find intensity of heritability significantly high until >30kb.



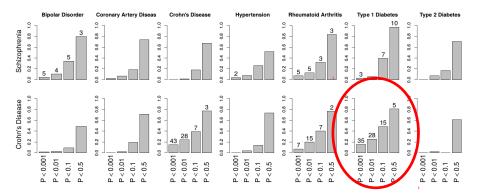
p-values for Schizophrenia and Crohn's obtained from independent studies.



SNPs associated with Crohn's are more important for Crohn's. (Good!)



SNPs associated with Schizophrenia are more important for Bipolar.



Find concordance between Crohn's and Type 1 Diabetes.

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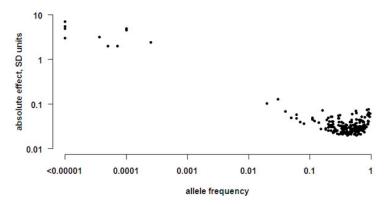
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Examining the Relationship between MAF and Effect Size

The default assumption is that all SNPs (GCTA) or genetic variations (LDAK) contribute equal heritability

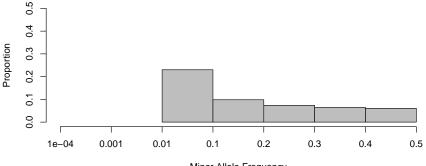
This seems to be supported by empirical evidence for some traits



Genetic architecture of body size in mammals, Kemper et al (2012)

(If true), this gives us an idea of how heritability varies with MAF

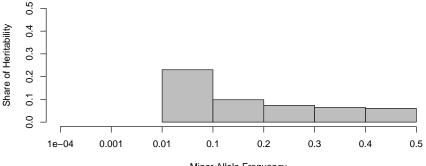
Distribution of MAF – COMMON SNPs



Minor Allele Frequency

(If true), this gives us an idea of how heritability varies with MAF

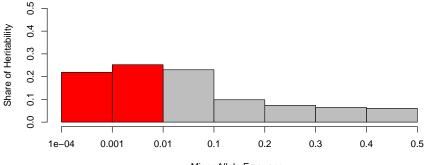
Distribution of Heritability – COMMON SNPs



Minor Allele Frequency

(If true), this gives us an idea of how heritability varies with MAF

Distribution of Heritability – ALL SNPs

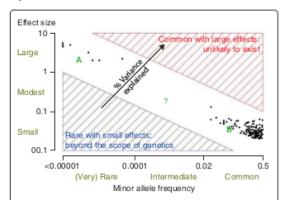


Minor Allele Frequency

In particular, it suggests the relative contribution of rare (red) and common (grey) variants

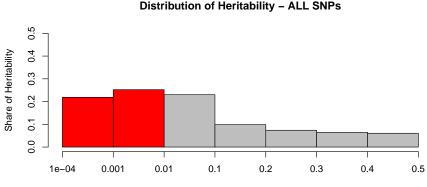
Is this Assumed Distribution Accurate?

Must recognise there is an implicit bias, because we tend to only find the highest heritability variants



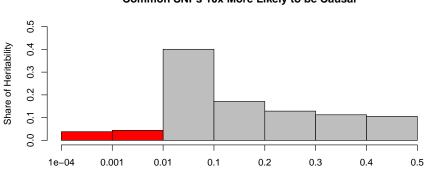
Understanding complex traits: from farmers to pharmas, Speed and Balding et al (2012)

Also, this assumes rare SNPs are equally likely to be causal as common SNPs. Is this a fair assumption?



Minor Allele Frequency

Things would change if the probability of a variant being causal depended on MAF



Common SNPs 10x More Likely to be Causal

Minor Allele Frequency

Testing How Heritability Varies with MAF

We can test this by performing heritability analysis with different kinship matrices, each corresponding to a different assumed relationship, and seeing which fare best

$$\mathcal{K}_{ij} = rac{1}{\sum_{w_j}} \sum w_j (X_{ij} - mean(X_j)) (X_{ij} - mean(X_j)) imes [Var(X_j)]^lpha$$

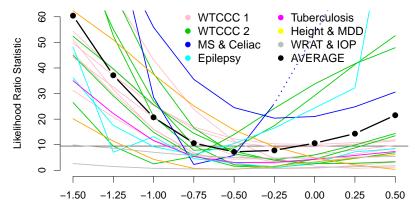
Smaller α means that rare variants contribute more h^2

The default in humans is $\alpha = -1$ (all variants contribute equal h^2)

The default in animals is $\alpha = 0$ (more common \Rightarrow more h^2)

We can now analyse our data using multiple α and see which fits best

Results for 22 GWAS Traits



Standardization Power

Smaller LRT (y-axis) \Rightarrow more support for corresponding α We deduce that $\alpha = -.5$ bests fit for these 22 traits

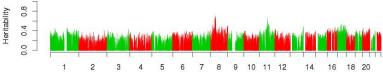
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Variance Explained by Copy Number Variants

We are not restricted to considering SNPs

the methodology can be applied to any data type

For breast cancer data, we constructed kinship matrices based on copy number variants (so now, each row of X represents the number of copies of a variant possessed by each individual)





Chromosomal Location of Gene Probe Probe

Instead of variance explained by SNPs, we estimate variance explained by copy number (where the phenotype is gene expression).

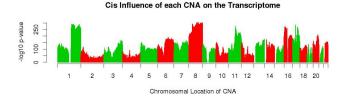
Variance Explained by Gene Expression

Alternatively, you can construct kinship matrices based on gene expressions (each row of X now represents intensity of a particular gene probe)

emper-double-1 2 3 4 5 6 7 8 9 10 11 12 14 16 18 20

Chromosomal Location of CNA

Genome-wide Influence of each CNA on the Transcriptome



We are now estimating variance explained by gene expression (where the phenotype is copy number change at a particular locus).

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Trait 1:
$$Y_1 = Z\alpha_1 + \beta_1 X_1 + \beta_2 X_2 + \ldots + \beta_{500\ 000} X_{500\ 000} + e_1$$

= $Z\alpha_1 + g_1 + e_1$

Trait 2:
$$Y_2 = Z\alpha_2 + \gamma_1 X_1 + \gamma_2 X_2 + \ldots + \gamma_{500\ 000} X_{500\ 000} + e_2$$

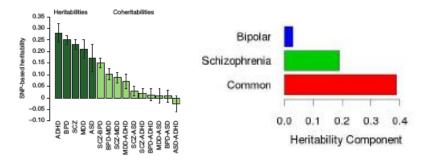
= $Z\alpha_2 + g_2 + e_2$

Now interested in the correlation between genetic effects: $\rho = cor(g_1, g_2)$.

Or equivalently can think of the average correlation between effect sizes: $\rho = cor(\beta_j, \gamma_j)$

Genetic relationship between five psychiatric disorders estimated from genome-wide SNPs

Cross-Disorder Group of the Psychiatric Genomics Consortium*



Hong will explain more in Module 17

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Introduced in Module 5

Suppose we want to test SNP j for association with Y

Standard Single-SNP Analysis:

$$Y = \beta_j X_j + e$$

Mixed Model Association Analysis:

$$Y = \beta_j X_j + g + e \quad g \sim \mathbb{N}(0, K\sigma_g^2)$$

Including the background (polygenic) random effect g effectively allows for confounding due to familiar relatedness and population structure

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Standard GWAS analysis examines each SNP individually.

There are many reasons we might prefer a gene-based analysis.

Reduces total number of tests Biologically plausible Can accumulate evidence across SNPs Standard GWAS analysis examines each SNP individually.

There are many reasons we might prefer a gene-based analysis.

Reduces total number of tests Biologically plausible Can accumulate evidence across SNPs

Our software GBAT performs set-based tests of association.

Fast and Powerful. Bayesian version accommodate prior information. Suitable for meta-analysis.

To estimate the total contribution of all SNPs, we use the model:

$$Y = \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3 + \beta_4 X_4 + \beta_5 X_5 + \beta_6 X_6 + \beta_7 X_7 + \beta_8 X_8 + \beta_9 X_9 + \beta_{10} X_{10} + \beta_{11} X_{11} + \beta_{12} X_{12} + \beta_{13} X_{13} + \beta_{14} X_{14} + \beta_{15} X_{15} + \beta_{16} X_{16} + \beta_{17} X_{17} + \beta_{18} X_{18} + \beta_{19} X_{19} + \beta_{20} X_{20} + \beta_{21} X_{21} + \beta_{22} X_{22} + \beta_{23} X_{23} + \beta_{24} X_{24} + \beta_{25} X_{25} + \beta_{26} X_{26} + \beta_{27} X_{27} + \beta_{28} X_{28} + \dots + \beta_{500\ 000} X_{500\ 000} + e.$$

To test a set of SNPs S, can reduce it to:

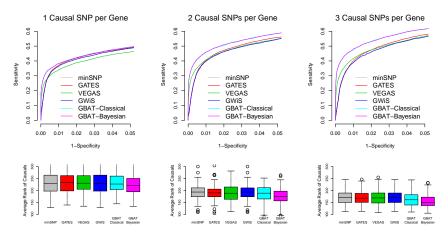
$$Y = \beta_{1}X_{1} + \beta_{2}X_{2} + \beta_{3}X_{3} + \beta_{4}X_{4} + \beta_{5}X_{5} + \beta_{6}X_{6} + \beta_{7}X_{7} + \beta_{8}X_{8} + \beta_{9}X_{9} + \beta_{10}X_{10} + \beta_{11}X_{11} + \beta_{12}X_{12} + \beta_{13}X_{13} + \beta_{14}X_{14} + \beta_{15}X_{15} + \beta_{16}X_{16} + \beta_{17}X_{17} + \beta_{18}X_{18} + \beta_{19}X_{19} + \beta_{20}X_{20} + \beta_{21}X_{21} + \beta_{22}X_{22} + \beta_{23}X_{23} + \beta_{24}X_{24} + \beta_{25}X_{25} + \beta_{26}X_{26} + \beta_{27}X_{27} + \beta_{28}X_{28} + \dots + \beta_{500\ 000}X_{500\ 000} + e.$$

i.e.,
$$Y = \sum_{j \in S} X_j \beta_j + e$$
 with $\beta_j \sim \mathbb{N}(0, \sigma_S^2/N_S)$.

Perform a likelihood ratio test for $\mathbb{P}(\sigma_{S}^{2} > 0)$.

Simulation Study

Generate phenotypes where 50 out of 1000 genes contribute heritability.



GBAT most powerful and fastest (does not require permutations).

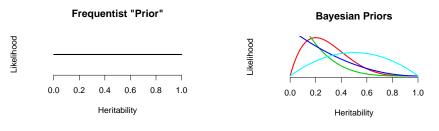
Frequentist methods base inferences only on evidence from data.

Bayesian methods combine evidence from data with prior beliefs.

Accommodating Prior Information

Frequentist methods base inferences only on evidence from data.

Bayesian methods combine evidence from data with prior beliefs.



Frequentist analysis assumes a gene is as likely to explain 1% or 99% of heritability.

Bayesian GBAT lets the user specify a prior distribution for heritability.

Single-SNP Meta-Analysis

Meta-analysis allows us to combine evidence for SNP associations across cohorts.

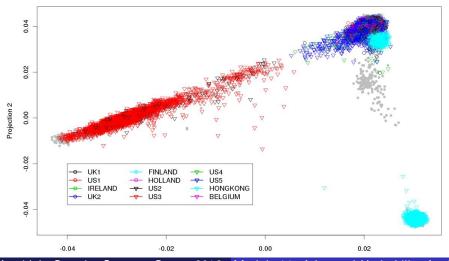
Study or sub-category	mean difference (SE	mean difference (random) 95% Cl	Weight %	mean difference (random) 95% Cl
Silver	6.9000 (3.1900)		→ 8.59	6.90 [0.65, 13.15]
Davas Pulse Cyc	7.5000 (3.0100)		→ 9.13	7.50 [1.60, 13.40]
Pakas High Pred	12.4000 (3.3600)		→ 8.10	12.40 [5.81, 18.99]
Pakas Low Pred	-0.7000 (1.9300)		13.12	-0.70 [-4.48, 3.08]
Hoyles	2.4000 (2.0800)		12.50	2.40 [-1.68, 6.48]
Nadashkevich	1.5000 (2.0800)		12.50	1.50 [-2.58, 5.58]
Tashkin	-1.0000 (0.0764)		18.86	-1.00 [-1.15, -0.85]
Airò	2.0000 (0.9100)		17.20	2.00 [0.22, 3.78]
Fotal (95% CI)		-	100.00	2.83 [0.35, 5.31]
fest for heterogeneity: Chi	² = 44.66, df = 7 (P < 0.00001), I	= 84.3%		
Test for overall effect: Z =	2.24 (P = 0.03)	190002000		

Existing gene tests are not well-suited for meta-analysis

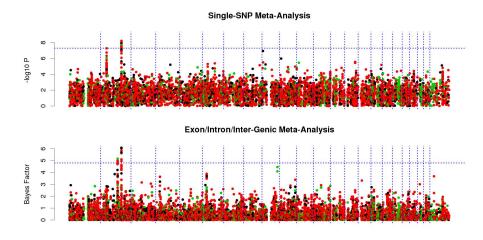
But when we score genes based on heritability, meta-analysis becomes straightforward

12-Cohort Gene-Based Meta-Analysis for Epilepsy

Work for the International League Against Epilepsy (ILAE)

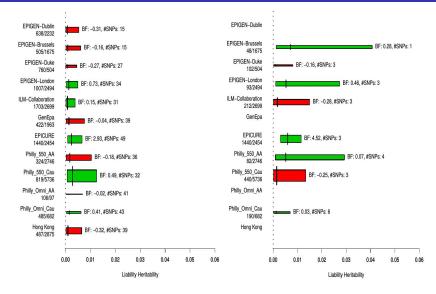


12-Cohort Gene-Based Meta-Analysis for Epilepsy



BLACK: EXONS — RED: INTRONS — GREEN: INTER-GENIC

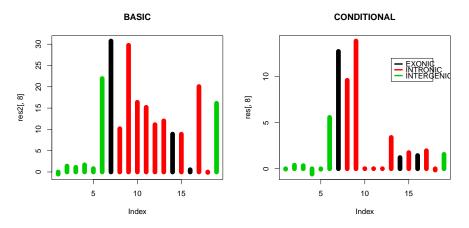
Two Hits on Chromosome 2



SCN1A

C17orf76-AS1

Fine Mapping of SCN1A



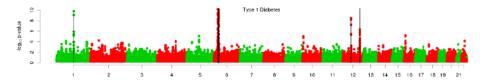
For each region of interest, can regress phenotype on each exon/intron/inter-genic chunk conditional on all other SNPs

- 1 Estimating SNP Heritability
- 2 Genome Partitioning
- 3 Intensity of Heritability
- 4 Testing Different Models
- 5 Using Other Datatypes
- 6 Bivariate Analysis
- 7 Mixed Model Analysis
- 8 Gene-Based Association Testing
- O Adaptive MultiBLUP
- 10 Applying to Animal and Plant Data

Adaptive MultiBLUP - See Module 11

Step 1: Divide genome into (say) 75kbp overlapping chunks.

Step 2: Test each chunk for association (using GBAT).



Step 3: Identify all significant chunks (say $P < 10^{-5}$). (Merge these chunks with neighbouring chunks with P < 0.01.)

E.g., for Type 1 Diabetes, obtain 4 local regions.

Step 4: Run MultiBLUP with five random effects.

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Applying these Methods to Animal Data

Human Genetics

Low Relatedness

10 000s of Individuals

Millions of SNPs

Binary Traits

Loadsa Money

Animal Genetics

High Relatedness

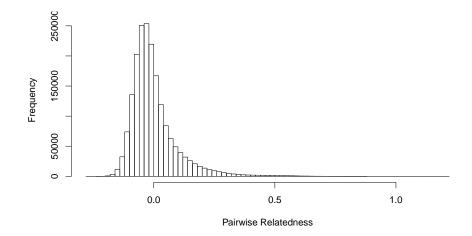
100/1000s of Individuals

10/100 000s of SNPs

Quantitative Traits

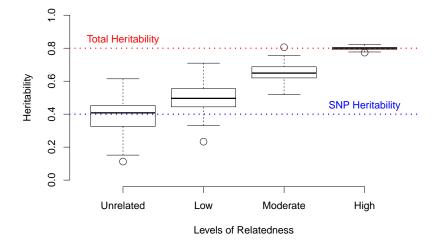
Nice Waterproof Clothing

Wellcome Trust Hetergeneous Stock Mice



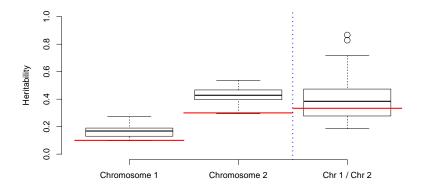
1940 mice, descended from 8 founders. 10091 SNPs.

Estimating SNP Heritability



WILL NOT WORK - Requires individuals to be unrelated

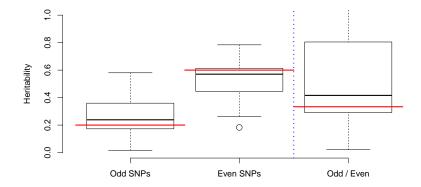
Simulated 80% heritability, of which 10% on Chr 1, 30% on Chr 2.



SEMI WORKS - Absolute values inflated, but relative values reasonable

Intensity of Heritability

SNPs 1, 3, 5, ..., 10091 explain 20% of variance SNPs 2, 4, 6, ..., 10090 explain 60% of variance



WORKS - Although precision low even with sparse SNP density

Mixed Model Association Analysis

Lellei

Nature Genetics **38**, 203 - 208 (2005) Published online: 25 December 2005 | <u>doi</u>:10.1038/ng1702

A unified mixed-model method for association mapping that accounts for multiple levels of relatedness

Jianming Yu^{1,2}, Gael Pressoir^{1,2}, William H Briggs², Irie Vroh Bi¹, Masanori Yamasaki³, John F Doebley², Michael D McMullen^{3,4}, Brandon S Gaut⁵, Dahlia M Nielsen⁶, James B Holland^{4,2}, Stephen Kresovich^{1,3},⁸ & Edward S Buckler^{1,4,3}

As population structure can result in spurious associations, it has constrained the use of association studies in human and plant genetics. Association mapping, however, holds great promise if true signals of functional association can be separated from the vast number of false signals generated by population structure^{1,2}. We have developed a unified mixed-model approach to account for multiple levels of relatedness simultaneously as detected by random genetic markers. We applied this new approach to two

WORKS - this is what it was designed for!

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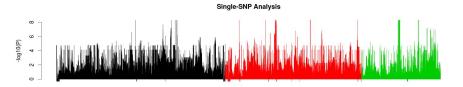
SEARCH PUBMED FOR

Data from Dan Jeffares

Collection of 161 natural isolates, phenotyped for 53 measurements of cell shape and growth rates.

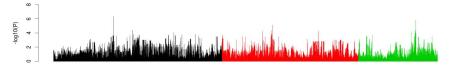
Data highly structured; while genotyped for over 1 000 000 SNPS, the effective size of the genome is closer to 1000.

Analysis of Schizosaccharomyces pombe



Standard Single-SNP analysis leads to high inflation





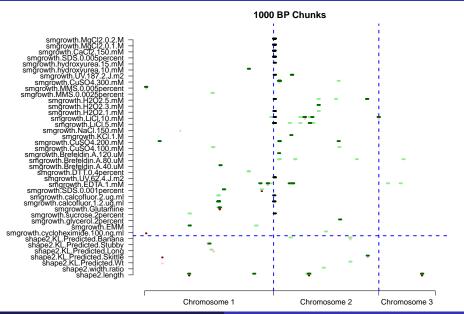
This is controlled by mixed model association analysis, but power is low

On next slide: across 58 traits, green are loci significant through chunk-based testing, red are those significant through single-SNP test

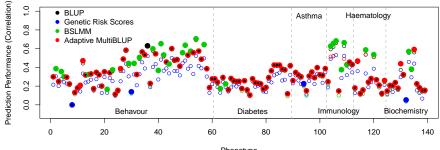
There are many more green (chunk-significant) hits than red (single-SNP-significant) hits

A gene (chunk) based test can have much higher power, because it considers groups of variants (typically in high LD) together and can afford a far lower significance threshold

Analysis of Schizosaccharomyces pombe



Adaptive MultiBLUP



Phenotype

WORKS - across 143 mice traits find modest improvement over BLUP

Yang, Visscher *et al.* showed that by applying mixed-models to GWAS data for unrelated individuals, it is possible to estimate h_{SNP}^2

While this was a major breakthrough in the missing heritability discussion, in my view, this is just the tip of the iceberg

By realising that different ${\bf K}$ correspond to different underlying models, we can now use (extensions of) SNP-based heritability analysis to investigate traits in fantastic detail

e.g., genome partitioning, intensity of heritability, distribution of h^2 , gene-based association analysis, MultiBLUP prediction

While estimating h_{SNP}^2 require unrelated individuals

MANY OF THE EXTENSIONS DO NOT