



Lecture 9: Introduction to Statistical Inference

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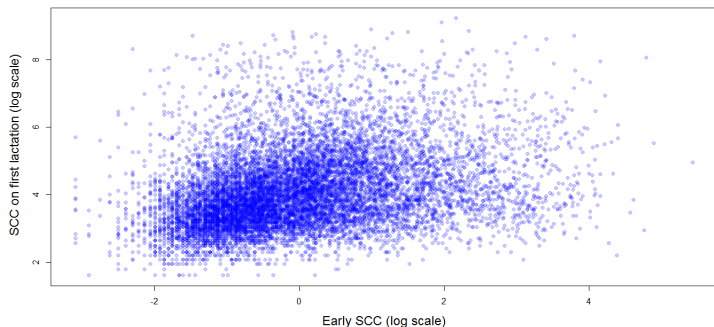


Overview

- Why do we need statistical models?
- The likelihood function
- Frequentist inference: maximum likelihood estimates and confidence intervals
- Likelihood and frequentist inference for stochastic SIR models

Why do we need models? An example

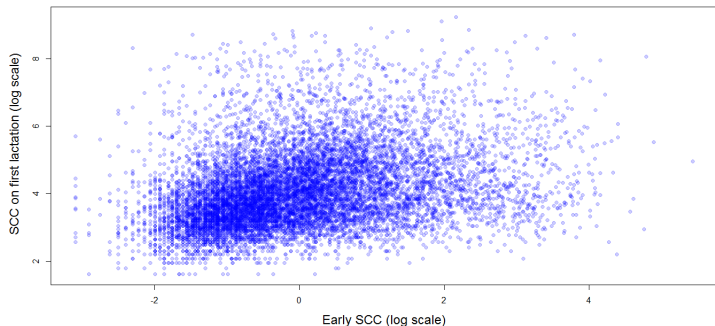
Early vs first day lactation somatic cell counts of young cows
(De Vliegher *et al.*, 2004)



- somatic cell count (SCC) is an indicator of milk quality and cow health
- early indicators of SCC can be useful to guide farm management

Somatic cell count (SCC) study: Research questions

Early vs first day lactation somatic cell counts of young cows



- how early SCC is related to first lactation SCC?
- is early SCC **really** related to first lactation SCC?
- are there any other variables associated with first lactation SCC??

models can provide answers to these questions

Models

Models are devices to **answer questions** and **represent reality**

mathematical models use **equations** to represent relationships

example: 1st lactation SCC = $\alpha + \beta(\text{early lactation SCC})$

- mathematical models represent **assumptions** and **underlying knowledge** about quantities of interest

problem: these models do not deal with the **uncertainty** regarding the phenomenon.

(is early SCC really related to 1st lactation SCC? are there any other variables associated with 1st lactation SCC?)

How to deal with the uncertainty underlying the problem?

Statistical models

Statistics deals with **uncertainty** by incorporating **variation** into the model

Sources of variation

- **systematic (deterministic) variation**: this can be based on **knowledge about the system** (example: early lactation SCC)
- **random (stochastic) variation**: this is due to **unknown factors/variables** which might be affecting the response

Statistics uses **probability distributions** to deal with random variation

example: linear regression model: $y = \alpha + \beta x + \epsilon$, $\epsilon \sim N(0, \sigma^2)$

ϵ has a normal distribution with mean 0 and variance σ^2

probability distributions represent data variation

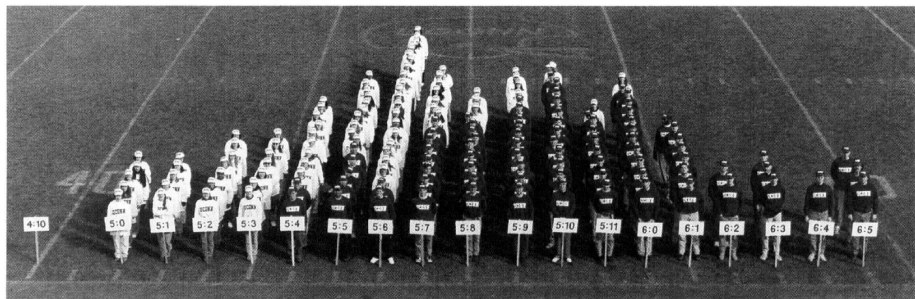
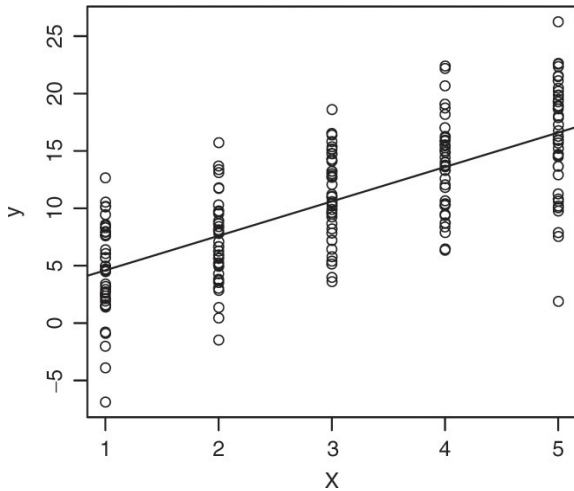


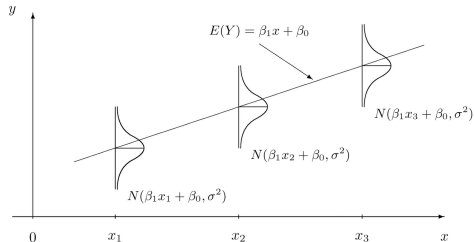
Figure 7. Living histogram of 143 student heights at University of Connecticut.

- What probability distribution gives the best fit to these data?
- Assumptions must be considered and evaluated based on available data

Variation in a variable might depend on the variation in another variable



Regression Analysis idea



linear regression model:

$$y = \beta_0 + \beta_1 x + \epsilon,$$
$$\epsilon \sim N(0, \sigma^2)$$

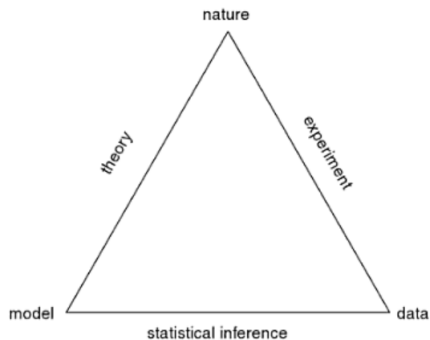
Distribution of the response variable Y given (fixed) x

$$Y|x \sim N(\beta_0 + \beta_1 x, \sigma^2)$$

task: estimate parameters β_0, β_1 based on a sample of the population: **Statistical Inference**

Parameter inference is one of the goals of Statistics

Scientific Method and Statistics



important statistical tasks

- design of experiments
- **inference**
- prediction/forecasting

statistical conclusions must be translated back into biology

Statistical inference

Statistical inference is the process of reaching **conclusions** from **data**

- data are always limited: usually a sample and/or limited experiments
- information may be limited even when dealing with large datasets (ex. gene expression data)
- different data provide different answers

any **statistical conclusion** involves a degree of **uncertainty**

Statistical inference tasks

- point estimation
- interval estimation
- hypothesis testing (ex. p-values, GWAS)

Point estimation

- Statistics uses **probability** to deal with random variation
- a probability distribution is assumed for a random variable of interest
- probability distributions are functions of **unknown parameters**

point estimation idea:

given the **available data** and assuming an **underlying model** for the variation observed (probability distribution) what single value is plausible for the indexing parameter?

a point estimate is the best guess about the parameter of a distribution

Maximum likelihood estimates (MLE)

Key idea (Fisher, 1922):

ML estimates **maximize** the probability of having observed the available data (e.g best explain the data)

Example: suppose an artificial data set with 3 **independent** observations:

$$x_1 = 3, x_2 = 4 \text{ and } x_3 = 8$$

Assume these data come from random variable X which follows a geometric distribution:

$$P(X = x; \theta) = (1 - \theta)^{x-1} \theta, \quad x = 1, 2, 3, \dots$$

The geometric distribution depends on an unknown parameter θ which can be estimated using x_1, x_2 and x_3 .

- In the example we assume that X_1, X_2, X_3 are **independent** and follow the **same distribution**
- If that's the case, X_1, X_2, \dots, X_N **independent** and **identically distributed (i.i.d)** random variables
- The assumption of i.i.d random variables is frequently considered in Statistics, **but it is not suitable for infectious disease data**

Example (cont'd)

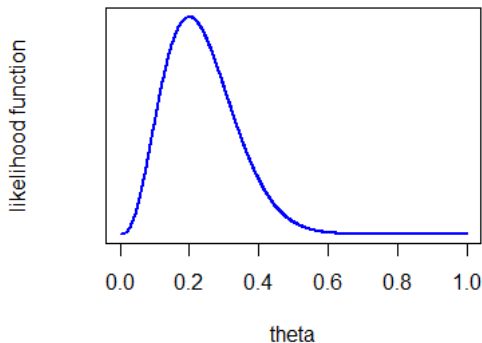
assuming i.i.d random variables (random sample), the probability of observing x_1, x_2 and x_3 is

$$\begin{aligned}P(X_1 = x_1, X_2 = x_2, X_3 = x_3; \theta) &= \\&= P(X_1 = 3; \theta)P(X_2 = 4; \theta)P(X_3 = 8; \theta) \\&= (1 - \theta)^{3-1}\theta(1 - \theta)^{4-1}\theta(1 - \theta)^{8-1}\theta \\&= (1 - \theta)^{12}\theta^3\end{aligned}$$

This is the **likelihood function**

the likelihood is a function of the unknown parameter θ .

Plotting the likelihood function



- the likelihood function is maximized at the value $\theta = 0.2$
- Hence, 0.2 is the maximum likelihood estimate of θ

Point estimation: another example

- Suppose an experiment to study the incidence of a certain tumour in mice
- a binary random variable (tumour/not tumour) can be used to represent each mouse in a random sample with size N
- the **total** number of mice with tumour can be modelled with a binomial distribution
- given that 6 out of 54 mice were observed with tumour, the probability of this event is

$$P(X = k) = \binom{54}{6} \theta^6 (1 - \theta)^{54-6}$$

where θ represents the unknown proportion of mice with tumour in the **population**. This is the likelihood function of θ for the data observed, and it is maximized at $\hat{\theta} = 0.11$

Hence, $\hat{\theta} = 0.11$ is the maximum likelihood estimate of the proportion of the mice with tumour in the population.

notes on point estimation/MLEs

- maximum likelihood estimators have important statistical properties (e.g., unbiasedness and efficiency)
- calculus or numeric procedures are used to obtain MLEs
- shape of the likelihood function plays a vital role in statistical inference (more on this later)
- it can be difficult to derive a proper likelihood to represent the data (example: epidemic modelling)
- **different samples provide different likelihoods hence different MLEs** - How to consider this feature into our inferences?

Two main approaches can be used: Frequentist (Classical) Inference and Bayesian Inference

Interval estimation: the frequentist approach

interval estimation idea:

given the available data and assuming an underlying model for the variation observed, what **range of values** is plausible for the indexing parameter?

This range of values involve a degree of uncertainty.

goal: for an **unknown** and **fixed** parameter θ , we want a range (a,b) such that

$$P[a < \theta < b] = \delta$$

- a and b are functions of the data: **a and b are random quantities**
- δ is a probability: usually 0.95
- example: when $\delta = 0.95$, the range of values $[a, b]$ is called a **95% confidence interval for the parameter θ**

Confidence intervals example: the frequentist approach

- Suppose that observations from a data set follow a normal distribution with unknown parameter μ and unknown variance σ^2
- X_1, X_2, \dots, X_N are random variables which represents the measurements of a random sample of size N
- the sample mean $\bar{X} = (X_1 + X_2 + \dots + X_N)/N$ is the **maximum likelihood estimator** of the population mean μ
- Since \bar{X} is a function of random variables, this **estimator** is also a **random variable** which follows a **probability distribution**

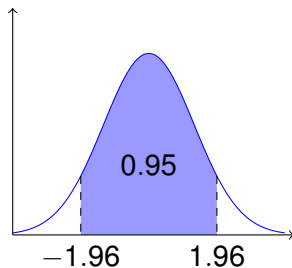
Confidence intervals example: the frequentist approach

- When the sample size is large, it can be shown \bar{X} has approximately a normal distribution unknown mean μ and variance $\hat{\sigma}^2/n$.
- $\hat{\sigma}^2$ is an estimate of the variance (the sample variance)
- this is a key result in Statistics (Central-limit theorem) and **it holds for any data distribution**

a 95% confidence interval for the
mean μ is

$$[\bar{X} - 1.96\hat{\sigma}/\sqrt{n}, \bar{X} + 1.96\hat{\sigma}/\sqrt{n}]$$

density of a $N(0,1)$



Interpretation of a confidence interval the frequentist approach

- Confidence intervals are functions of the data available: they are random quantities
- In the frequentist approach, the distribution parameters are assumed fixed

frequentist interpretation of a 95% CI for the mean μ :

If a large number of confidence intervals were calculated using independent random samples of the population, 95% of them would contain the true mean μ

notes on frequentist interval estimation

for normal distributed variables

- exact confidence intervals can be obtained even when the sample size is small: in this case a Student's t-distribution is used
- confidence intervals for the variance can be calculated using a chi-square distribution

Other distributions

- large-sample confidence intervals can be used to calculate CI for proportions (practical)
- definition of large-sample can vary. Usually $n=30$ is enough for means of normal distributions and also proportions

point and interval estimation can be also applied to regression analysis

Summary of key ideas: Frequentist inference

- Statistics allow the incorporation of uncertainty about the quantities of interest into models
- Any statistical conclusion involves uncertainty
- Frequentist inference allows parameter estimation by using available data only (using likelihood)
- Maximum likelihood estimates are the ones that best explain the observed data
- confidence intervals represent the uncertainty regarding Frequentist inference: plausible **range of values** for a model parameter

Tutorial 9a

Analysing fish infection data

- Looking at likelihood plots
- Comparing frequentist estimates based on different sample sizes

Representing infectious disease dynamics: stochastic compartmental models

“diversity” represented through disease states - ex. SIR model



possible individual states

- susceptible (S)
 - infected (I)
 - recovered (R)
- Infections occur with rate $\beta S(t)I(t)$ (Poisson process)
 - Recoveries occur with rate $\gamma \rightarrow$ infectious period follows an exponential distribution (general stochastic model)

Assumptions

- closed population, homogeneous mixing and no latent period
- all individuals are **equally** susceptible and infectious

Inference problem: estimation of β , γ and R_0

Estimating R_0 (in general)

There are several approaches for estimating R_0 , depending on **assumptions** and **data limitations** (see Keeling and Rohani's book)

For example, R_0 can be estimated using reported cases, seroprevalence data, Average age at infection and final size data

most of simple methods for estimating R_0 are based on deterministic models - cannot accommodate uncertainty regarding parameter estimation

R_0 can be also estimated from transmission experiments (see Diekmann *et al*, 2013) - **practical**

Likelihood function for a SIR model

The likelihood function depends on the **structure** and **availability** of the epidemic data

Some possible scenarios:

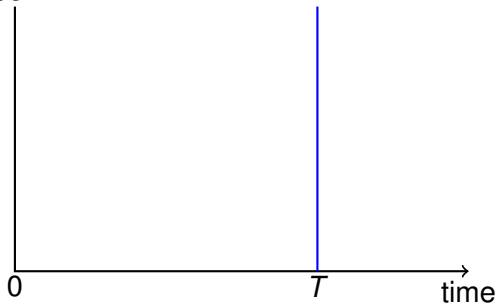
- **best scenario:** exact infection and removed times (complete observation)
 - Infection times observed at fixed time periods (e.g an individual infected between week 1 and 2)
 - only removed times observed
 - only final numbers infected
- } **Not suitable for genetic analysis**

Infectious disease data in genetic analysis

usual phenotype: binary disease status at a **single time** T

animal phenotype

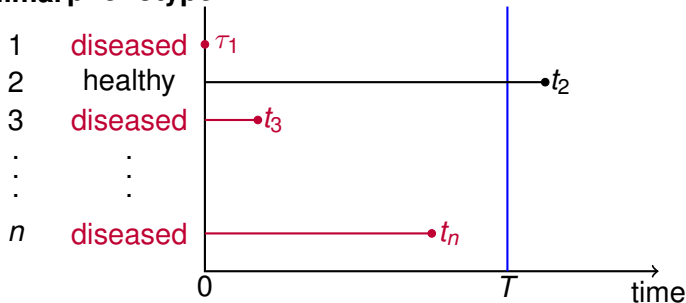
1	diseased
2	healthy
3	diseased
⋮	⋮
⋮	⋮
n	diseased



Infectious disease data in genetic analysis

usual phenotype: binary disease status at a **single time** T

animal phenotype

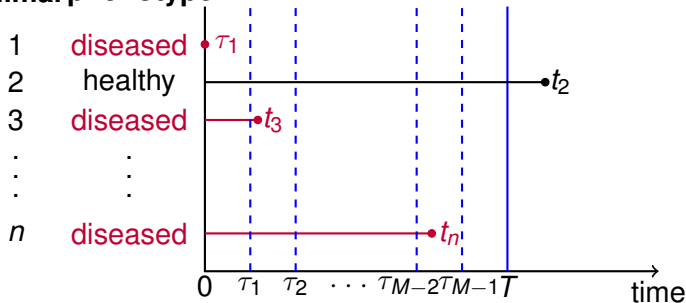


- time to infection is more informative about disease traits (rarely observed but can be inferred - to be seen in the MCMC lecture)

Infectious disease data in genetic analysis

usual phenotype: binary disease status at a **single time** T

animal phenotype

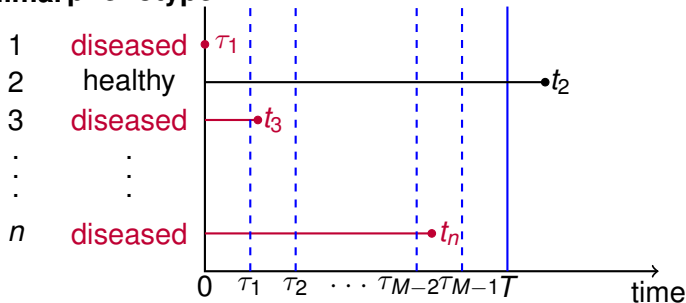


- time to infection is more informative about disease traits (rarely observed but can be inferred - to be seen in the MCMC lecture)
- disease status at sampling times ($0, \tau_1, \tau_2, \dots, T$) are required

Infectious disease data in genetic analysis

usual phenotype: binary disease status at a **single time** T

animal phenotype



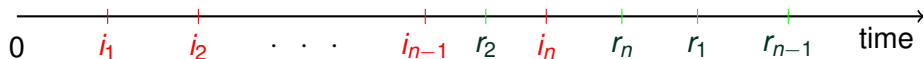
- time to infection is more informative about disease traits (**rarely observed but can be inferred - to be seen in the MCMC lecture**)
- disease status at sampling times ($0, \tau_1, \tau_2, \dots, T$) are required
- better genetic analyses of disease traits when using binary longitudinal data (e.g, Anacleto *et al*, 2015)

Likelihood function for a SIR model (idea)

assumption: Complete observation of the data

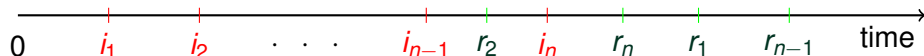
- infection times: $\mathbf{i} = (i_2, i_3, \dots, i_n)$
- removal times: $\mathbf{r} = (r_1, r_2, \dots, r_n)$

data observed until the end of the epidemic:



- **infection events are not independent!**
- likelihood can be decomposed into the contributions of the infectious and the removal processes
- order of the events and special properties of the stochastic process and associated distributions are important for derivation of the likelihood

Likelihood function for a SIR model (idea)



likelihood contribution of the infectious process

- Time between infections follows an exponential distribution (Poisson Process property)
- based on the conditional densities at small time steps

likelihood contribution of the removal processes

- based on the exponential distribution of the infectious periods
- infectious period of individual $j = r_j - i_j$

Different likelihood derivations depending on inference approach

- See Bailey and Thomas (1971), Becker (1989) and Britton and O'Neill (2002). See also Kypraios (thesis, 2007) for comparison.

Maximum likelihood estimates of SIR model parameters

Assume that

- epidemic started with an initially infected individual and was observed until its end at time T
- n_I and n_R are the total numbers of infecteds and recovered
- S_t and I_t and R_t are the **numbers** of susceptibles, infecteds and recovered at time t

$$\hat{\beta} = \frac{n_I}{\int_{I_1}^T S_t I_t dt}$$

$\hat{\beta}$'s denominator is the accumulated rate of contacts between susceptibles and infecteds

$$\hat{\gamma} = \frac{n_R}{\int_{I_1}^T R_t dt}$$

$\hat{\gamma}$'s denominator is the aggregated length of the infectious period

Maximum likelihood estimates of SIR model parameters

Rida (1991) showed that standard errors of $\hat{\beta}$ and $\hat{\gamma}$ are

$$\text{s.e}(\hat{\beta}) = \frac{\hat{\beta}}{\sqrt{n_l - 1}} \quad \text{and} \quad \text{s.e}(\hat{\gamma}) = \frac{\hat{\gamma}}{\sqrt{n_r}}$$

Hence, approximate confidence intervals for β and γ can be obtained by normal approximation

Approximate confidence intervals can also be derived for R_0 (see Diekmann *et al*, 2013)

Summary of key ideas: Frequentist inference for SIR model parameters

- challenges: high dependence between infection events and incomplete data
- several approaches for estimating R_0
- likelihood function of the SIR model is derived depending on inference approach
- likelihood function of the SIR can be decomposed into the contributions of the infectious and the removal processes - order of events are important!
- individual level data are required for (most of) genetic analysis of infectious diseases

References

Statistical Inference

- Wasserman, Larry. All of statistics: a concise course in statistical inference. Springer, 2013.

Frequentist Inference for stochastic epidemic models

- Becker, N. G., Britton, T. (1999). Statistical studies of infectious disease incidence. Journal of the Royal Statistical Society: Series B (Statistical Methodology), 61(2), 287-307.
- Diekmann, Odo, Hans Heesterbeek, and Tom Britton. Mathematical tools for understanding infectious disease dynamics. Princeton University Press, 2012.
- Britton, T. (1998). Estimation in multitype epidemics. Journal of the Royal Statistical Society: Series B (Statistical Methodology), 60(4), 663-679.

Guidelines for Statistical Practice

- Kass, R.E., Caffo, B., Davidian, M., Meng, X.-L., Yu, B., and Reid, N. (2016) Ten simple rules for effective statistical practice, PLoS Computational Biology.

Tutorial 9b

- Frequentist inference for R_0
- Comparing frequentist estimates based on different sample sizes