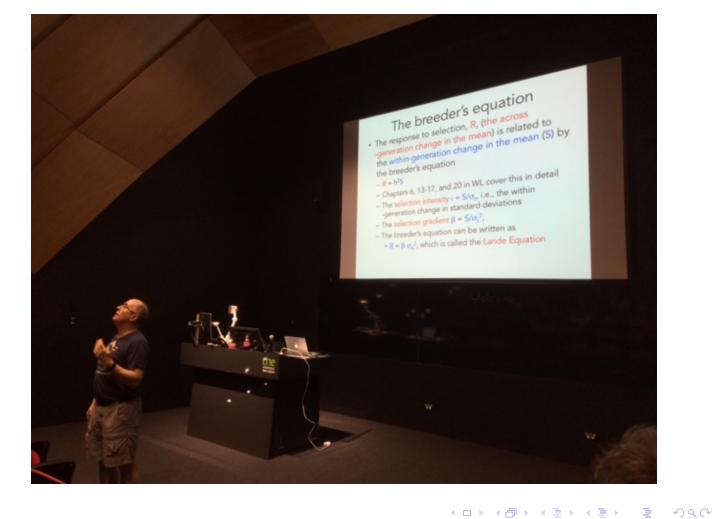
## Analysis of univariate phenotypic selection

Michael Morrissey February 3, 2020

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#### ► My goals

 Key concepts in methods and theory to support solid empirical work

# Michael Morrissey Analysis of univariate phenotypic selection Preliminaries 1

- My goals
  - Key concepts in methods and theory to support solid empirical work

#### Structure

- This lecture: jump right in! quite fine detail for a "simple" case
- ▶ Subsequent lectures: elaboration of simple univariate case

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- ► My goals
  - Key concepts in methods and theory to support solid empirical work
- Structure
  - This lecture: jump right in! quite fine detail for a "simple" case
  - ▶ Subsequent lectures: elaboration of simple univariate case
- References
  - Very few on slides
  - Online slides with notes have extensive and specific references to W&L2018
  - See notes for other references as well as notes about unpublished results



In addition to Julius and Bruce, I am able to be here thanks to:









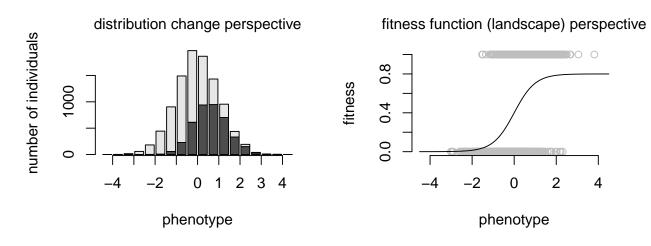
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- $\blacktriangleright$  z: phenotype
- $\blacktriangleright$  a: breeding value
- $\blacktriangleright$  W: (absolute) fitness
- ▶ w: relative fitness  $(w_i = \frac{W_i}{W})$
- $\blacktriangleright \bar{x}, \mu_x, E[x]$ : mean of x
- $\triangleright$   $V_x, \sigma_x^2, VAR[x]$ : variance of x
- $\triangleright \sigma_{x,y}, COV[x,y]$ : covariance of x and y
- $\blacktriangleright \beta_{xy}, b_{y|x}: \text{ regression of } y \text{ on } x$



Two complimentary ways of thinking about natural selection:



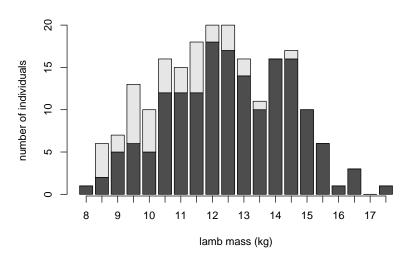
Key concepts to look out for in each framework

selection differentials the breeder's equation selection gradients the Lande equation

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#### A natural summary of selection:



- mean mass before selection:  $\mu_0 = 12.35 \text{ kg}$
- ▶ mean mass after selection:  $\mu_1 = 12.74 \text{ kg}$
- change in mass:  $S = \mu_1 \mu_0 = 0.39$  kg



Justification comes from the mechanics of evolution

$$evolution = f(genetics, selection)$$

For S, the justification is this:

$$R = h^2 S$$

Interpretation of  $h^2$ :

- ▶ ratio of heritable to total variance  $\frac{V_a}{V_p}$
- ▶ slope of the parent-offspring regression  $b_{o|mp}$

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By construction, the regression of offspring phenotype on mid-parent phenotype is a function that predicts offspring phenotype according to

 $z_o = \mu + b_{o|mp}(z_{mp} - \mu) + e$ 

Where a definition of  $h^2$  is  $h^2 = b_{o|mp}$ .

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#### Derivation of the Breeder's equation

By construction, the regression of offspring phenotype on mid-parent phenotype is a function that predicts offspring phenotype according to

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Where a definition of  $h^2$  is  $h^2 = b_{o|mp}$ . The expectation of a linear transformation of a random variable x with expectation E[x], according to the transformation y = a + bx is E[y] = a + bE[x], so

$$E[z_o] = \mu + b_{o|mp} E[z_{mp} - \mu]$$

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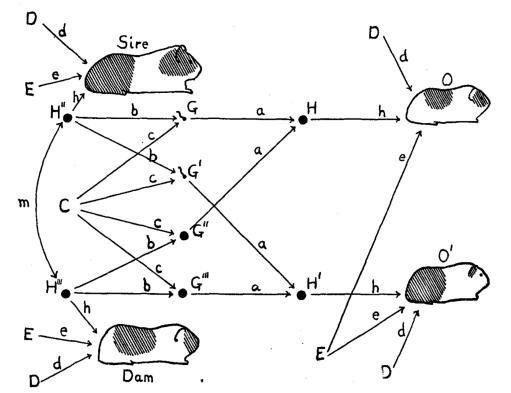
 $\mu$  is not a random variable insofar as our analysis is concerned, so

$$E[z_o] - \mu = b_{o|mp}(E[z_{mp}] - \mu)$$
$$R = E[z_o] - \mu = b_{o|mp}S$$

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Michael Morrissey Analysis of univariate phenotypic selection Where does  $h^2$  come from?

$$h^2 + d^2 + e^2 = 1$$



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Suppose a population contains n individuals, indexed i, with individual fitness  $W_i$ . For e.g.,  $W_i = 0$  if dead,  $W_i = 1$  if alive:

$$S = \mu_{after} - \mu_{before}$$
  
=  $\frac{1}{n} \sum_{i} \frac{W_{i}}{\bar{W}} z_{i} - \frac{1}{n} \sum_{i} z_{i}$   
=  $E[wz] - (1)E[z]$  (with  $w = W\bar{W}^{-1}$  such that  $\bar{w} = 1$ )  
=  $COV[z, w]$ 

Notes:

▶ this is a proof of the Robertson-Price identity

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 this allows calculation of S as the mean weighted by relative fitness, for fitness components other than viability

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#### Background on OLS regression

If the covariance of A and B is  $\sigma_{AB}$  and the variance of A is  $\sigma_A^2$ then the regression of B on A is given by

$$\beta_{AB} = \frac{\sigma_{AB}}{\sigma_A^2}$$

This, or its multivariate equivalent, is exactly what your favourite software does to give you regression coefficients.

For multiple regression, if  $\Sigma_{\mathbf{x}}$  is the covariance matrix of the predictor variables, and  $\Sigma_{\mathbf{x}y}$  is a (column) vector of covariances of predictors with the response, then the gradient of partial regression coefficients is

$$\boldsymbol{\beta} = \boldsymbol{\Sigma}_{\mathbf{x}}^{-1} \boldsymbol{\Sigma}_{\mathbf{x}y}$$

The univariate case is key to the next slide, the multivariate case comes up in multivariate selection.

Recall that

$$R = h^2 S$$

and that

$$h^2 = \frac{V_a}{V_p}$$

 $\mathbf{SO}$ 

$$R = \frac{V_a}{V_p}S = V_a \frac{S}{V_p}$$

recall also that S = COV[z, w], so

$$R = V_a \beta_{zw}$$

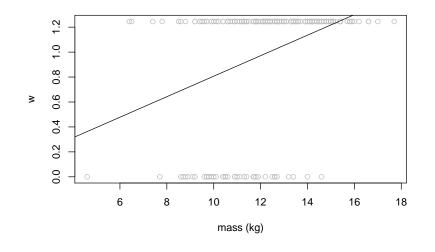
This is the univariate Lande equation.



#### Using regression to estimate $\rho$

Recall that we previously considered the mean of survivors relative to the unselected mean to calculate S.

The same data could have been plotted as a scatter plot, making regression natural.



$$\beta = 0.082$$

units are  $kg^{-1}$  because w is unitless

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Since

$$\beta = \frac{S}{V_z}$$

rearrangement yields

$$S = V_z \beta$$

In ewe lambs  $V_z$  of mass is 4.78, and  $\beta = 0.082$ , so

$$S = 4.78 \cdot 0.082 = 0.39$$

which is exactly what we got for S in the first place.



Is selection of S = 0.5 kg of lamb mass stronger or weaker than (also positive directional) selection of S = 50 mm of oak tree sapling height?

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#### Standardisations of differentials and gradients 1: $\sigma$

Is selection of S = 0.5 kg of lamb mass stronger or weaker than (also positive directional) selection of S = 50 mm of oak tree sapling height? Some kind of standardisation is required for most comparisons of selection coefficients.

- Standardising to unit variance is by far the most common in empirical studies.
- $\blacktriangleright$  variance-standardising S:

$$S_{\sigma} = \frac{S}{\sigma_z}$$

• variance-standardising  $\beta$ :

$$\beta_{\sigma} = \beta \cdot \sigma_z$$

• recall that 
$$\beta = \frac{S}{\sigma_z^2}$$
, so

$$\beta_{\sigma} = \frac{S}{\sigma_z^2} \cdot \sigma_z = \frac{S}{\sigma_z} = S_{\sigma}$$

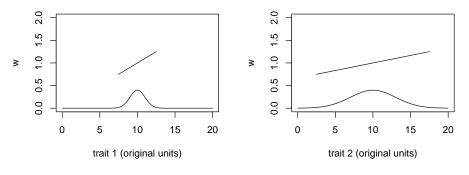
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#### Properties of $S_{\sigma}$ and $\beta_{\sigma}$

Consider these two associations between a trait and relative fitness:



Which is stronger selection?

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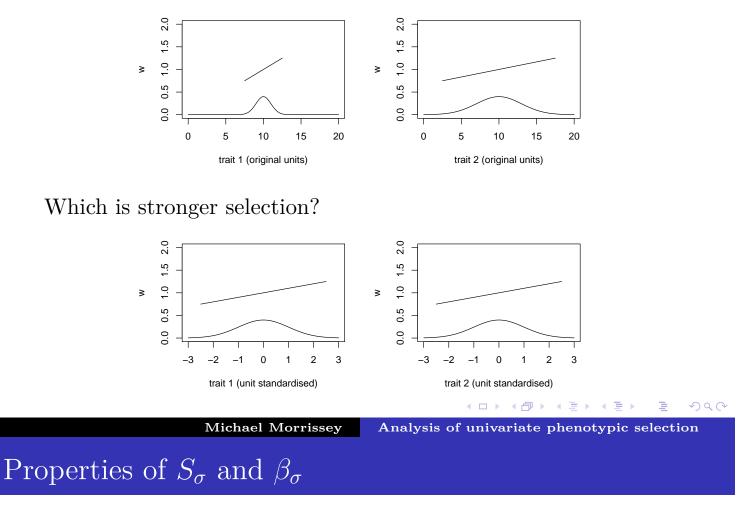
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#### Properties of $S_{\sigma}$ and $\beta_{\sigma}$

Consider these two associations between a trait and relative fitness:



- ►  $\beta_{\sigma}$  (or  $S_{\sigma}$ ) gives the slope of the relative fitness function (or change in mean phenotype), in units of phenotypic standard deviations.
- ► These are not all-purpose measures the strength of selection
- A shallow function (low β in original units), can cause a lot of variation in expected fitness, if there is a big range of phenotype (also in its original units)
- ▶  $\beta_{\sigma}$  and  $S_{\sigma}$  are the standard deviation of relative fitness implied by the trait-fitness association
  - this confounds (not necessarily in a pejorative sense) phenotypic variability and steepness of the effect of phenotype on fitness

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mean-standardising S:

$$S_{\mu} = \frac{S}{\mu_z}$$

Q: By what percent are survivors larger (smaller) than the initial average? A:  $S_{\mu}(\cdot 100)$ . mean-standardising  $\beta$ :

$$\beta_{\mu} = \beta \cdot \mu_z$$

Q: By what percent does a 1% change in phenotype change relative fitness? A:  $\beta_{\mu}$ .

There is no direct equivalence between  $S_{\mu}$  and  $\beta_{\mu}$ , as there is for  $S_{\sigma}$  and  $\beta_{\sigma}$ 

$$\beta_{\mu} = \beta \cdot \mu_{z} = \frac{S}{\sigma_{z}^{2}} \cdot \mu_{z}$$
$$= \frac{S_{\mu}\mu_{z}}{\sigma_{z}^{2}} \cdot \mu_{z} = S_{\mu}\frac{\mu_{z}^{2}}{\sigma_{z}^{2}}$$

Michael Morrissey Analysis of univariate phenotypic selection Evolvability and mean standardisation

In terms of the mean, how much evolution do we expect?

$$\begin{aligned} \Delta \bar{z} &= V_a \beta \\ \frac{\Delta \bar{z}}{\bar{z}} &= \frac{V_a \beta}{\bar{z}} \\ \frac{\Delta \bar{z}}{\bar{z}} &= \frac{V_a \frac{\beta_\mu}{\bar{z}}}{\bar{z}} \\ \frac{\Delta \bar{z}}{\bar{z}} &= \frac{V_a}{\bar{z}^2} \beta_\mu \end{aligned}$$

 $\frac{V_a}{\bar{z}^2}$  has been termed the *evolvability*, and is closely related (and referred to essentially interchangeably with the *coefficient of additive genetic variance*  $CV_a = \frac{\sigma_a}{\mu}$ .

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#### $h^2$ and alternative standardisations of S

▶  $h^2$  is a variance-standardisation of the genetic variability in a population

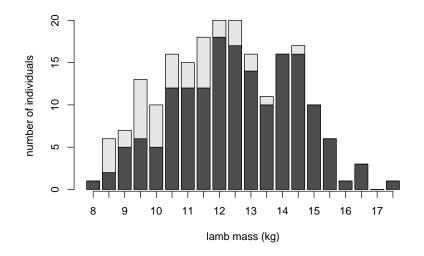
 the breeder's equation holds, using h<sup>2</sup>, for any standardisation of traits

$$R = h^2 S$$
$$R_{\sigma} = h^2 S_{\sigma}$$
$$R_{\mu} = h^2 S_{\mu}$$

as we saw on the previous slide, for the Lande equation to hold, V<sub>a</sub> must be expressed in the same standardising unit (e.g., σ or μ) in which the gradient and response are expressed.



Selection may change the variability of a population



$$\sigma_0^2 = 4.78$$

$$\sigma_1^2 = 4.38$$

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Care is needed: purely directional selection changes the variance too:

 $\Delta \sigma_z^2(directional) = -S^2$ 

So, the change in the variance, over and above the effect of purely directional selection to reduce the variance, could be defined as

$$C = \Delta \sigma_z^2 + S^2$$

In ewe lambs:

• 
$$\mu_0 = 12.35, \, \mu_1 = 12.74, \, \text{so} \, S = 0.39$$
  
•  $\sigma_0^2 = 4.78, \, \sigma_1^2 = 4.38$ 

So,

$$\Delta \sigma_z^2 = 4.38 - 4.78 = -0.40$$

and

$$C = -0.40 + 0.40^2 = -0.25$$

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Michael Morrissey Analysis of univariate phenotypic selection The Lande-Arnold regression

Just like the change in the mean is related to a linear regression, the change in the variance is related to a quadratic regression coefficient.

Lande and Arnold (1983) showed that

$$w_i = \alpha + \beta (z_i - \bar{z}) + \gamma \frac{1}{2} (z_i - \bar{z})^2 + e_i$$

and that when the phenotype is Gaussian,

$$C = \gamma \cdot \sigma_z^4$$

(note  $\sigma_z^4 = \left(\sigma_z^2\right)^2$ ) and

$$\Delta \sigma_z^2 = \sigma_z^4 \left( \gamma - \beta^2 \right)$$

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$$w_i = \alpha + \beta (z_i - \bar{z}) + \gamma \frac{1}{2} (z_i - \bar{z})^2 + e_i$$

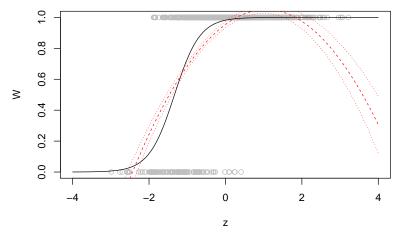
- does not (and neither does OLS, regardless of what the textbooks say) assume normality of residuals
- does assume normality of phenotype (in quadratic case), despite this not generally being an assumption of OLS
- heterogeneity of residual variance does affect OLS SEs (but no effect on estimates), but this is probably a minimal concern
- $\blacktriangleright$  calculation of w is surprisingly frequently messed up
- mean-centering is critical in quadratic case

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• factor of 1/2 is very easy to miss

#### Some further notes about $\gamma$

Consider this relationship between trait and fitness:



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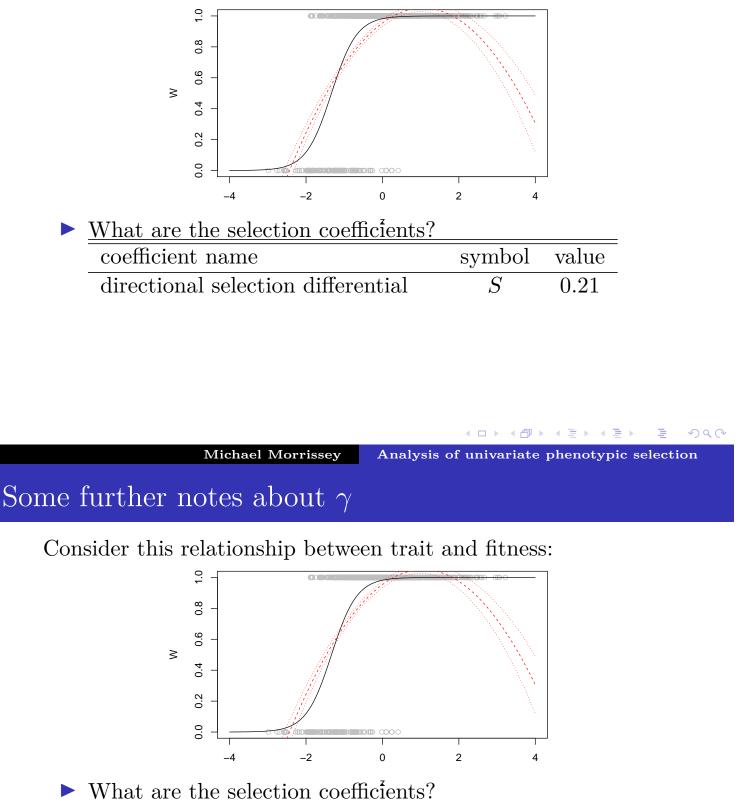
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#### Some further notes about $\gamma$

Consider this relationship between trait and fitness:



coefficient name	symbol	value
directional selection differential	S	0.21
directional selection gradient	eta	0.21

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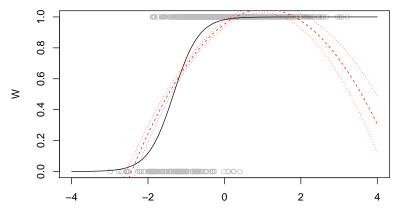
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#### Some further notes about $\gamma$

Consider this relationship between trait and fitness:



What are the selection coefficients?		
coefficient name	symbol	value
directional selection differential	S	0.21
directional selection gradient	eta	0.21
change in the phenotypic variance	$\Delta \sigma_z^2$	-0.24

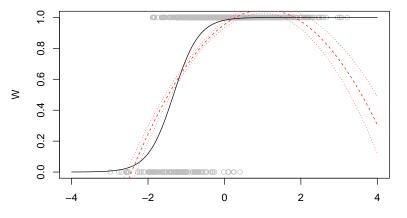
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#### Some further notes about $\gamma$

Consider this relationship between trait and fitness:



#### • What are the selection coefficients?

coefficient name	$\operatorname{symbol}$	value
directional selection differential	S	0.21
directional selection gradient	eta	0.21
change in the phenotypic variance	$\Delta \sigma_z^2$	-0.24
stabilising selection differential	C	-0.20

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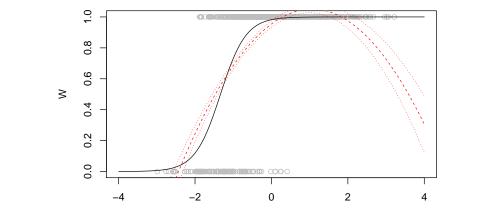
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coefficient name	symbol	value
directional selection differential	S	0.21
directional selection gradient	eta	0.21
change in the phenotypic variance	$\Delta \sigma_z^2$	-0.24
stabilising selection differential	C	-0.20
quadratic selection gradient	$\gamma$	-0.20
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#### Michael Morrissey Analysis of univariate phenotypic selection Fitness "functions" and fitness "landscapes" 1

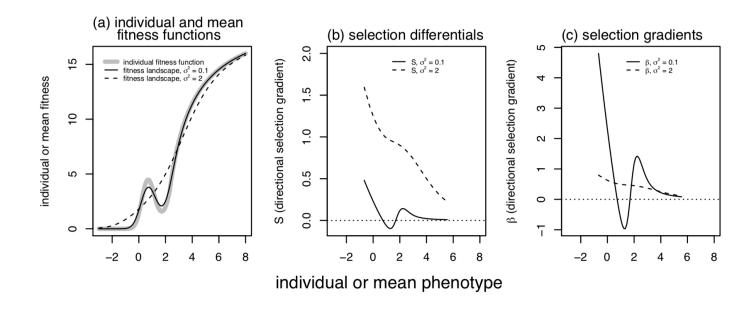
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- The main message from the previous slide is that selection coefficients represent very specific things about natural selection, they are not catch-all representations of trait-fitness relationships
- Directional and quadratic gradients can be thought of as the average slope and curvature of the of a fitness function, in the region of phenotype in a population.
- As such differentials and gradients reflect not only the ecological relationship between trait and fitness, but also the distribution of phenotype along the x-axis of the function mapping trait on to fitness.

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A super-handy result from Charles Stein (1973) is that if y = f(x), then

$$COV[x, y] = VAR[x]E[f'(x)]$$

So, if W = f(z)

$$S \cdot \bar{W} = COV[z, W] = \sigma_z^2 E[f'(z)]$$
  
$$\beta = \sigma_z^2 E[f'(z)] \bar{W}^{-1}$$

So, for any arbitrary function, we can calculate a selection gradient.

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### A few notes about $\beta_{average}$ gradient

- If z is normal, then  $\beta_{OLS} = \beta_{average \ gradient}$
- If z is not normal, then β, calculated as the average gradient, still works in the Lande equation, provided that breeding values are normal and uncorrelated with environmental effects. To see this, note that

$$COV[a,w] = V_a E[\frac{dz}{da}f'(z)]$$

and that  $\frac{dz}{da} = 1$ , so the change in breeding values (from applying the Robertson-Price identity to breeding values from one generation to the next

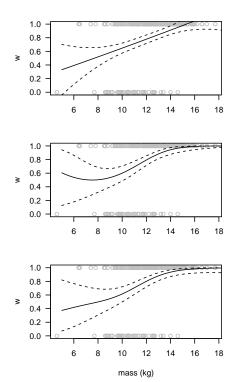
- $E\left[\frac{\partial w}{\partial z}\right] = \frac{1}{\bar{W}}\frac{\partial \bar{W}}{\partial \bar{z}}$  if changes in  $\bar{z}$  are understood to arise from only the mean changing.
  - this is useful for numerical implementation, and also should allow analysis of discontinuous fitness functions.



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#### Inference of selection gradients from arbitrary functions

The relationship between the average partial derivatives of the fitness function and selection gradients suggests a numerical scheme applicable to any fitness function shape



linear model (with quadratic term)  $\beta = 0.082, \ \beta_{\sigma} = 0.180$  $\gamma = 0.001, \ \gamma_{\sigma} = 0.003$ 

logistic regression model (with quadratic term)  $\beta = 0.080, \ \beta_{\sigma} = 0.176$  $\gamma = -0.0076, \ \gamma_{\sigma} = -0.033$ 

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generalised spline regression  $\beta = 0.079, \ \beta_{\sigma} = 0.173$  $\gamma = -0.010, \ \gamma_{\sigma} = -0.044$ 

- Some functions have direct relationships to selection coefficients
- Average derivative methods could be used to brute-force gradient calculations for shape of fitness function
- Analytical relations are still often very useful, especially for theory or predicting consequences of management
- No way that I can explain all the following equations, or that you can remember them. My purpose is to make you aware of the range of known relationships
- Useful type of relationship without analytical results: logistic and probit functions

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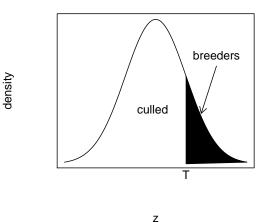
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Selection coefficients and other fitness functions - truncation

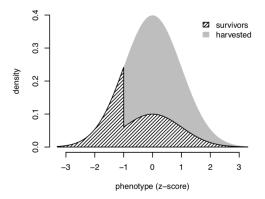


$$S = \sigma_z \frac{f_N(t)}{F_N - t}$$

where  $f_N()$  and  $F_N(t)$  are unit normal density and cumulative functions, and  $g = \frac{T - \mu_z}{\sigma_z}$ 

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Selection coefficients and other fitness functions - partial truncation



like truncation, but individuals above or below the critical trait value are culled with probability  $\alpha$ 

$$S = \sigma_z \frac{\alpha f_N(t)}{\alpha F_N - t - 1}$$

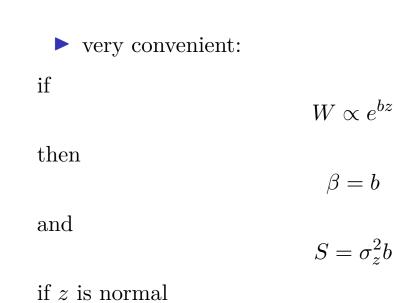
partial truncation can behave surprisingly differently to truncation selection!



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exponential



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# Selection coefficients and other fitness functions - Gaussian

If z is normal with mean  $\mu_z$  and variance  $\sigma_z^2$ , and

$$W(z) \propto e^{rac{-(z- heta)^2}{2\omega^2}}$$

then

 $\beta = -\mathcal{S}(\bar{z} - \theta),$ 

where  $S = \frac{1}{\sigma^2 + \omega^2}$  (regrettably, S is not the selection differential)

Michael Morrissey Analysis of univariate phenotypic selection

Selection coefficients and other fitness functions log-exponential (generalisation of gaussian)

Consider the fitness function

$$E[W(z_i)] = exp^{a+bz_i+\frac{1}{2}gz_i}$$

▶ looks an awful lot like a Lande-Arnold regression
 ▶ this is a Gaussian fitness function when g < 0 provided that g < <sup>1</sup>/<sub>σ<sup>2</sup><sub>z</sub></sub>

$$\beta = \frac{b + g\mu_z}{1 - g\sigma_z^2}$$

and

$$\gamma = \frac{b^2 + g(1-g)}{(1-g)^2}$$

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